

One-Way Analysis of Variance (ANOVA) is a method for comparing the means of a populations.

This kind of problem arises in two different settings

1. When a independent random samples are drawn from a populations.
2. When the effects of a different treatments on a homogeneous group of experimental units is studied, the group of experimental units is subdivided into a subgroups and one treatment is applied to each subgroup. The a subgroups are then viewed as independent random samples from a populations.

Notation:

| Group | Population | | Sample | | |
|-------|------------|----------|--------|-------------|-------|
| | Mean | SD | Size | Mean | SD |
| 1 | μ_1 | σ | n_1 | \bar{x}_1 | s_1 |
| 2 | μ_2 | σ | n_2 | \bar{x}_2 | s_2 |
| | | ... | | | |
| a | μ_a | σ | n_a | \bar{x}_a | s_a |

The hypotheses of interest in One-Way ANOVA are:

$$H_0 : \mu_1 = \mu_2 = \dots = \mu_a$$

$$H_A : \mu_i \neq \mu_j \text{ for some } i \neq j$$

Assumptions required for One-Way ANOVA

1. Random samples are independently selected from a (treatments) populations.
2. The a populations are approximately normally distributed.
3. All a population variances are equal.

The summary statistics and assumptions are the same assumptions as we made for the pooled t-test to compare two normal means, except that now we have $a \geq 2$ populations.

Notation/terminology

a is the number of factor levels (treatments) or populations

x_{ij} is the j th observation in the i th sample, $j = 1, \dots, n_i$

n_i is the sample size of the i th sample

$\bar{x}_{i.} = \sum_{j=1}^{n_i} x_{ij} / n_i$ is the i th sample mean

$s_i^2 = \frac{1}{(n_i-1)} \sum_{j=1}^{n_i} (x_{ij} - \bar{x}_{i.})^2$ is the i th sample variance

$\bar{x}_{..} = \frac{1}{n} \sum_{i=1}^a n_i \bar{x}_{i.}$ is the overall mean of all observations

$n = \sum_{i=1}^a n_i$ is the total number of observations

Sums of squares and degrees of freedom

- The total variability in the response is called the total sum of squares, SST.
- The total variability SST is partitioned into between treatment and within treatment sums of squares.
- The notations SS_{Tr} (treatment sum of squares) and SSB (between sum of squares) are synonymous.
- The notations SS_{Error} and SSE (error sum of squares) and SSW (within sum of squares) are synonymous.
- Following are the formulas for the **sums of squares**.

$$SST = \sum_{i=1}^a \sum_{j=1}^{n_i} (x_{ij} - \bar{x}_{..})^2$$

$$SS_{Tr} = \sum_{i=1}^a \sum_{j=1}^{n_i} (\bar{x}_{i.} - \bar{x}_{..})^2 = \sum_{i=1}^a n_i (\bar{x}_{i.} - \bar{x}_{..})^2$$

$$SSE = \sum_{i=1}^a \sum_{j=1}^{n_i} (x_{ij} - \bar{x}_{i.})^2 = \sum_{i=1}^a (n_i - 1) s_i^2$$

- Associated with each sum of squares is its degrees of freedom.
 - The **total degrees of freedom** is $n - 1$.
 - The **treatment degrees of freedom** is $a - 1$.
 - The **error degrees of freedom** is $n - a$.
 - There is an additivity relationship in the sums of squares.

$$SST = SS_{Tr} + SSE$$

- There is an additivity relationship for the degrees of freedom.

$$n - 1 = (a - 1) + (n - a)$$

- Mean squares: $MSE = SSE/(n - a)$, $MS_{Tr} = SS_{Tr}/(a - 1)$
- Observed value of test statistic: $F_{obs} = MS_{Tr}/MSE$
-

$$p - value = P(F_{a-1, n-a} \geq F_{obs})$$

Mean squares, F and p-values

Scaled versions of the treatment and error sums of squares (the sums of squares divided by their associated degrees of freedom) are known as **mean squares**: $MS_{Tr} = SS_{Tr}/(a - 1)$ and $MSE = SSE/(n - a)$.

- MS_{Tr} and MSE are both estimates of the error variance, σ^2 . MSE is always unbiased (its mean equals σ^2), while MS_{Tr} is unbiased only when the null hypothesis is true. When the alternative H_A is true, MS_{Tr} will tend to be larger than MSE .
- The ratio of the mean squares is $F = MS_{Tr}/MSE$. This should be close to 1 when H_0 is true, while large values of F provide evidence against H_0 . The null hypothesis H_0 is rejected for large values of the observed test statistic F_{obs} .
- The **p-value** is the probability that an F random variable with $a - 1$ numerator and $n - a$ denominator degrees of freedom is at least as large as F_{obs} , that is

$$p - value = P(F_{a-1, n-a} \geq F_{obs})$$

Calculations are conveniently displayed in an **ANOVA table**, as follows.

| Source | df | SS | MS | F_{obs} | p-value |
|------------|---------|-----------|-----------|-----------------------|--------------------------------|
| Treatments | $a - 1$ | SS_{Tr} | MS_{Tr} | $\frac{MS_{Tr}}{MSE}$ | $P[F_{a-1, n-a} \geq F_{obs}]$ |
| Error | $n - a$ | SSE | MSE | | |
| Total | $n - 1$ | SST | | | |

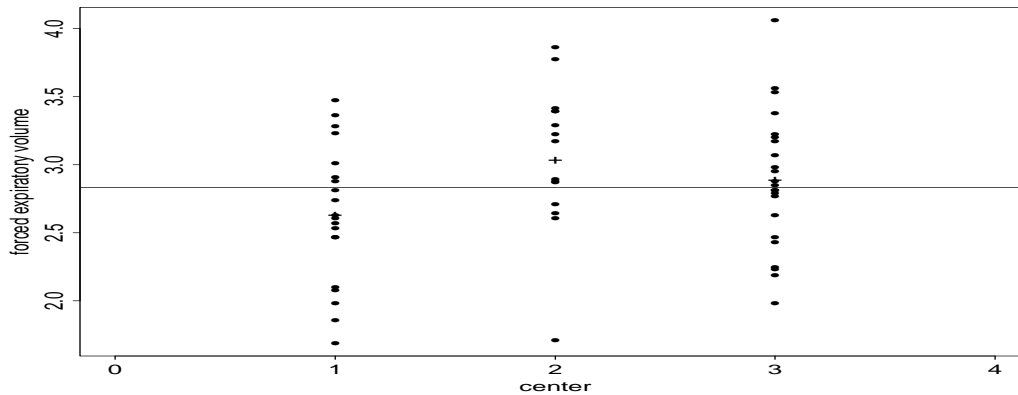
Review of F Distribution - using the F table

1. What is the probability that an F variable with 3 numerator and 5 denominator degrees of freedom is greater than 12.5? From the F table we see that $P(F_{3,5} > 12.06) = .01$ and $P(F_{3,5} > 33.20) = .001$, so that $.001 < P(F_{3,5} > 12.5) < .01$

Example: Pagano and Gauvreau gives the forced expiratory volume in 1 second for patients with coronary artery disease at three different centers.

| | Johns Hopkins | Rancho Los Amigos | St. Louis | Overall |
|-----------------|---------------|-------------------|-----------|---------|
| | 3.23 | 3.22 | 2.79 | |
| | 3.47 | 2.88 | 3.22 | |
| | 1.86 | 1.71 | 2.25 | |
| | 2.47 | 2.89 | 2.98 | |
| | 3.01 | 3.77 | 2.47 | |
| | 1.69 | 3.29 | 2.77 | |
| | 2.10 | 3.39 | 2.95 | |
| | 2.81 | 3.86 | 3.56 | |
| | 3.28 | 2.64 | 2.88 | |
| | 3.36 | 2.71 | 2.63 | |
| | 2.61 | 2.71 | 3.38 | |
| | 2.91 | 3.41 | 3.07 | |
| | 1.98 | 2.87 | 2.81 | |
| | 2.57 | 2.61 | 3.17 | |
| | 2.08 | 3.39 | 2.23 | |
| | 2.47 | 3.17 | 2.19 | |
| | 2.47 | | 4.06 | |
| | 2.74 | | 1.98 | |
| | 2.88 | | 2.81 | |
| | 2.63 | | 2.85 | |
| | 2.53 | | 2.43 | |
| | | | 3.20 | |
| | | | 3.53 | |
| n_i | 21 | 16 | 23 | 60 |
| $\sum y_{ij}$ | 55.15 | 48.52 | 66.21 | 169.88 |
| $\sum y_{ij}^2$ | 149.7581 | 151.2436 | 196.0483 | 497.05 |
| \bar{y}_i | 2.63 | 3.03 | 2.88 | 2.831 |
| s_i | 0.496 | 0.523 | 0.498 | .522 |
| SS_i | 4.924 | 4.107 | 5.450 | 16.063 |

Does the mean FEV differ in the three groups?



- $H_0 : \mu_1 = \mu_2 = \mu_3$
 $H_A : \text{at least two of } \mu_1, \mu_2, \mu_3 \text{ are different}$
- Using calculations as above, can find $SSE = 14.479$, $SST = 16.063$.
- Then complete the ANOVA table, starting with $SS_{Tr} = 16.063 - 14.481 = 1.582$
- The completed ANOVA table is:

| Source | Sum of Squares | DF | Mean Square | F |
|---------|----------------|----|-------------|------|
| Between | 1.582 | 2 | 0.791 | 3.11 |
| Within | 14.481 | 57 | .254 | |
| Total | 16.063 | 59 | | |

- degrees of freedom are 2 and 57. 57 is not in the table, so **go to next smaller df** = 40, and $p - \text{value} = P(F_{2,57} \geq 3.11) \approx P(F_{2,40} \geq 3.11) \in (.05, .1)$.
- Conclusion: reject H_0 at level α if and only $\alpha \geq .1$.

Following calculations were done using the R program, starting with summary statistics. Any differences from the numbers quoted above are due to rounding errors in the summary statistics.

```
> ni=c(21,16,23)           sample sizes
> ybari=c(2.63,3.03,2.88)  sample means
> si=c(.496,.523,.498)    sample standard deviations
> SSE=sum((ni-1)*si^2)     calculation of SSE
> SSE
[1] 14.47934
> ybar=sum(ybari*ni)/sum(ni)  calculation of overall mean
> ybar
[1] 2.8325                   overall mean - note the rounding error
> SStr=sum(ni*(ybari-ybar)^2) calculation of SStr
> SStr
[1] 1.537125                SStr has quite a bit of rounding error
```

A bit more detail on the underlying calculations

- The Total Sum of Squares is the sum of squared deviations from the overall average $SST = \sum \sum (y_{ij} - \bar{y})^2$ (generally done on computer)

$$(3.23 - 2.831)^2 + \dots + (3.53 - 2.831)^2 = 16.063$$

- The Within Sum of Squares is the sum of squared deviations from the group averages $SSE = \sum \sum (y_{ij} - \bar{y}_i)^2$ (done on computer)

$$(3.23 - 2.63)^2 + \dots + (3.22 - 3.03)^2 + \dots + (3.53 - 2.88)^2 \\ = 4.924 + 4.107 + 5.450 = 14.481$$

also calculated as a weighted sum of group variances $SSW = \sum (n_i - 1)s_i^2$ (easy calculation on hand calculator)

$$20(.496)^2 + 15(.523)^2 + 22(.498)^2 = 14.479$$

or as the sum of the within group sums of squares $SSW = \sum SS_i$

$$4.924 + 4.107 + 5.45 = 14.481.$$

- The Between Sum of Squares is the difference, $SSB = TSS - SSW$

$$16.063 - 14.481 = 1.582$$

also calculated as the weighted sum of squares of difference between group averages and the overall average $SSB = \sum n_i(\bar{y}_i - \bar{y})^2$ (easy calculation on hand calculator)

$$= 21(2.63 - 2.831)^2 + \dots + 23(2.88 - 2.831)^2 = 1.537$$

(differences are due to round-off error).

- The degrees of freedom (DF) are the number of independent pieces of information.

- For Between, it is one less than the number of groups, $a - 1$.
- For Within, it is one less than the number in each group, summed over groups, which is the same as the sample size less the number of groups $n - a$.
- For Total, it is one less than the total number of observations, $n - 1$.

The Test

- If the means are not different, SSB will be a small component of the total, i.e. small relative to SSW .
- Under the stated assumptions

$$F = \frac{SSB/(a - 1)}{SSW/(n - a)} = 3.11$$

has an F distribution with $a - 1$ and $n - a$ degrees of freedom.

- Values of F near 1 or smaller indicate no difference.
- Tables give only selected quantiles of F for selected degrees of freedom.
- When the dfs we want aren't in the tables, we use the next smaller degrees of freedom.
- Because 3.11 is between 2.44 and 3.23, the .10 and .05 quantiles of the F with 2 and 40 degrees of freedom, we conclude P is between .05 and .10.
- The computer gives .052.

Which means are different. There are 3 possible comparisons. If we do 3 t-tests, each with probability of type I error α , then the probability of committing at least one type I error is greater than α . To control the overall probability of type I error at α , we can use the Bonferroni procedure, as follows.

- If we determine there are differences among the groups, we want to identify them, usually by testing each pair, $H_0 : \mu_i = \mu_l$ versus $H_a : \mu_i \neq \mu_l$.
- To stop the overall error rate from growing, we decrease the error rate on each test.
- The Bonferroni correction uses $\alpha^* = \alpha/c$ for each test, where $c = \binom{a}{2}$ is the number of possible comparisons.

- We use the pooled estimate of standard deviation, $s = \sqrt{MSE}$, and the test statistic

$$t_{i,k} = \frac{\bar{x}_i - \bar{x}_k}{s \sqrt{\frac{1}{n_i} + \frac{1}{n_k}}}$$

which has a t distribution with $n - a$ degrees of freedom.

- We find significant evidence against $H_0 : \mu_i = \mu_k$ at level α if $p\text{-value} \leq \alpha^*$.
- In the example, if we used $\alpha = .10$ we would conclude there were significant differences among the groups using the F test.
- $t_{1,2} = -2.39$, $t_{1,3} = -1.64$ and $t_{2,3} = .91$ have p-values .02, .11 and .37 (using the computer), only the first is less than $\alpha^* = .10/3 = .033$.