# ACSC/STAT 4703, Actuarial Models II 

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## Midterm Examination <br> Model Solutions

Here are some values of the Gamma distribution function with $\theta=1$ that may be needed for this examination:

| $x$ | $\alpha$ | $F(x)$ | $x$ | $\alpha$ | $F(x)$ | $x$ | $\alpha$ | $F(x)$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 245 | 255 | 0.2697208 | 2.5 | 4 | 0.2424239 | 4.375 | 4 | 0.6361773 |
| $\left(\frac{7.5}{12}\right)^{3}$ | $\frac{4}{3}$ | 0.1117140 | 3.841 | 2.4 | 0.8409823 | 4.875 | 4 | 0.7169870 |
| $\left(\frac{9.5}{12}\right)^{3}$ | $\frac{4}{3}$ | 0.2507382 | 4.375 | 3 | 0.8118663 | 5.375 | 4 | 0.7837292 |
| 1.356 | 2.4 | 0.2801616 | 4.875 | 3 | 0.8644174 | 2.156 | 5 | 0.06782354 |
| 1.941 | 2.4 | 0.4612472 | 5.375 | 3 | 0.9035828 | 3.203 | 5 | 0.219922 |
| 2.367 | 2.4 | 0.5775816 | 3.875 | 4 | 0.5417358 | 8.542 | 5 | 0.9274742 |

Here are the critical values for a chi-squared distribution:

| Degrees of | Significance level |  |  |
| :--- | ---: | ---: | ---: |
| Freedom | $90 \%$ | $95 \%$ | $99 \%$ |
| 1 | 2.705543 | 3.841459 | 6.634897 |
| 2 | 4.605170 | 5.991465 | 9.210340 |
| 3 | 6.251389 | 7.814728 | 11.344867 |
| 4 | 7.779440 | 9.487729 | 13.276704 |
| 5 | 9.236357 | 11.070498 | 15.086272 |

1. Using an arithmetic distribution $(h=1)$ to approximate an inverse Weibull distribution with $\tau=2$ and $\theta=6$, calculate the probability that the value is more than 8.5, for the approximation using the method of local moment matching, matching 1 moment on each interval.
[Hint:

$$
\int_{8}^{9} x^{-2} e^{-\left(\frac{6}{x}\right)^{2}} d x=0.00840944
$$

I
If we let $r_{n}$ and $s_{n}$ be the probability assigned to $n$ from the intervals $[n-1, n]$ and $[n, n+1]$ respectively, so that $p_{n}=r_{n}+s_{n}$, then we have $p_{1}+\ldots+p_{7}+r_{8}=F_{X}(8)=e^{-\left(\frac{6}{8}\right)^{2}}=0.569782824731$. We have

$$
\begin{aligned}
s_{8}+r_{9} & =e^{-\left(\frac{6}{9}\right)^{2}}-e^{-\left(\frac{6}{8}\right)^{2}}=0.071397563699 \\
8 s_{8}+9 r_{9} & =\int_{8}^{9} 2 \frac{6^{2}}{x^{2}} e^{-\left(\frac{6}{x}\right)^{2}} d x=72 \times 0.00840944=0.60547968 \\
s_{8} & =9 \times 0.071397563699-0.60547968 \\
& =0.037098393291
\end{aligned}
$$

Thus $P\left(X_{a}>8.5\right)=1-0.569782824731-0.037098393291=0.393118781978$.
2. Claim frequency follows a negative binomial distribution with $r=3.3$ and $\beta=1.5$. Claim severity (in thousands) has the following distribution:

| Severity | Probability |
| ---: | :--- |
| 0 | 0.46 |
| 1 | 0.33 |
| 2 | 0.16 |
| $\geqslant 3$ | 0.05 |

The expected claim severity per loss is 0.81 . The company buys excess-of loss reinsurance for aggregate losses exceeding 2.
(a) Use the recursive method to calculate the probability that the reininsurance makes a payment.

For the negative binomial distribution, we have $a=\frac{\beta}{1+\beta}=0.6$ and $b=(r-1) \frac{\beta}{1+\beta}=1.38$. The recurrence relation is therefore

$$
f_{S}(x)=\frac{\sum_{y=1}^{x}\left(0.6+1.38 \frac{y}{x}\right) f_{X}(y) f_{S}(x-y)}{1-0.6 f_{X}(0)}
$$

We also have

$$
f_{S}(0)=P_{N}\left(f_{X}(0)\right)=(2.5-1.5 \times 0.46)^{-3.3}=0.141143441599
$$

We calculate

$$
\begin{aligned}
f_{S}(1) & =\frac{1.98 \times 0.33 \times 0.141143441599}{0.724}=0.127380006548 \\
f_{S}(2) & =\frac{1.29 \times 0.33 \times 0.127380006548+1.98 \times 0.16 \times 0.141143441599}{0.724}=0.136657335754
\end{aligned}
$$

Thus the probability that the reinsurance company makes a payment is $1-0.141143441599-0.127380006548-$ $0.136657335754=0.594819216099$.
(b) What is the expected payment on the reinsurance? [Hint: first calculate the insurer's expected payment with this reinsurance policy. Then consider the expected total payments between the insurer and the reinsurer.]

The insurer pays 0 if $S=0,1$ if $S=1$ and 2 if $S \geqslant 2$. Thus the insurer's expected payment is $f_{S}(1)+2\left(1-f_{S}(0)-\right.$ $\left.f_{S}(1)\right)=2-0.127380006548-2 \times 0.141143441599=1.59033311025$.
The expected number of losses is $3.3 \times 1.5=4.95$, and the expected claim per loss is 0.81 , so the expected aggregate claim is $4.95 \times 0.81=4.0095$. The insurer pays 1.59033311025 of this, so the reinsurer's expected payment is $4.0095-1.59033311025=2.41916688975$.
3. An insurance company collects a sample of 700 claims. Based on previous experience, it believes these claims might follow an inverse Pareto distribution with $\theta=0.7$ and $\tau=3.9$. To test this, it computes the following p-p plot.

(a) How many of the claims in their sample were between 3 and 11?

We have that

$$
\begin{aligned}
F^{*}(3) & =\left(\frac{3}{3+0.7}\right)^{3.9}=0.441353087998 \\
F^{*}(11) & =\left(\frac{11}{11+0.7}\right)^{3.9}=0.786152158728
\end{aligned}
$$

From the graph:

we read that $F_{n}(3) \approx 0.76$ and $F_{n}(11) \approx 0.98$. Since there are 700 samples, there are approximately $(0.98-0.76) \times$ $700=154$ samples between 3 and 11 in the dataset.
[There are actually 174 samples between 3 and 11 in the dataset.]
(b) Which of the following is a plot of $D(x)=F_{n}(x)-F^{*}(x)$ for this data?


Justify your answer.
From the p-p plot, we see that $F^{*}(x)<F_{n}(x)$ for all $x$, so $D(x)$ should be positive for all $x$. This is only the case for (iii). Therefore (iii) is the true plot of $D(x)$.
4. An insurance company collects the following sample:

$$
\begin{array}{lllllllllllllll}
0.06 & 0.32 & 0.61 & 0.67 & 1.16 & 2.53 & 4.02 & 5.09 & 10.27 & 15.83 & 17.64 & 17.84 & 20.00 & 20.92 & 24.44
\end{array} 42.52
$$ 63.8071 .84

They model this as following a distribution with the following distribution function:

| $x$ | $F(x)$ | $i^{2}\left(\log \left(F\left(x_{i+1}\right)\right)-\log \left(F\left(x_{i}\right)\right)\right)$ | $(18-i)^{2}\left(\log \left(1-F\left(x_{i}\right)\right)-\log \left(1-F\left(x_{i+1}\right)\right)\right)$ |
| :---: | :--- | :---: | :---: |
| 0.06 | 0.003940665 | 3.6581388 | 52.466840 |
| 0.32 | 0.152854776 | 1.3462645 | 21.656703 |
| 0.61 | 0.214016777 | 0.8161002 | 6.703294 |
| 0.67 | 0.234330430 | 5.0055286 | 26.831221 |
| 1.16 | 0.320402414 | 1.5140605 | 5.855879 |
| 2.53 | 0.340406390 | 4.2591188 | 11.325024 |
| 4.02 | 0.383158609 | 13.8026392 | 32.511770 |
| 5.09 | 0.507824310 | 1.0051731 | 1.992625 |
| 10.27 | 0.515863074 | 2.8523340 | 3.893844 |
| 15.83 | 0.534352306 | 9.9014547 | 10.302913 |
| 17.64 | 0.589968955 | 12.5596075 | 10.958980 |
| 17.84 | 0.654497867 | 17.0641094 | 13.339362 |
| 20.00 | 0.736838758 | 13.0086484 | 9.131373 |
| 20.92 | 0.795796371 | 5.5330811 | 2.957822 |
| 24.44 | 0.818581810 | 22.8721986 | 10.549221 |
| 42.52 | 0.906170573 | 4.0967235 | 1.524215 |
| 63.80 | 0.920788516 | 6.6446792 | 1.260842 |
| 71.84 | 0.942204512 | 19.2886673 |  |
|  |  | 145.228528 | 223.261928 |

Calculate the Anderson-Darling statistic for this model and this data.
The Anderson-Darling statistic for this data is

$$
\begin{aligned}
& -n F^{*}(u)+n \sum_{i=1}^{k}\left(F_{n}\left(y_{i}\right)\right)^{2}\left(\log \left(F^{*}\left(y_{i+1}\right)\right)-\log \left(F^{*}\left(y_{i}\right)\right)+n \sum_{i=0}^{k}\left(1-F_{n}\left(y_{i}\right)\right)^{2}\left(\log \left(1-F^{*}\left(y_{i}\right)\right)-\log \left(1-F^{*}\left(y_{i+1}\right)\right)\right.\right. \\
& =n\left(\sum _ { i = 1 } ^ { k } \frac { i ^ { 2 } } { n ^ { 2 } } \left(\log \left(F^{*}\left(y_{i+1}\right)\right)-\log \left(F^{*}\left(y_{i}\right)\right)+\sum_{i=0}^{k} \frac{(n-i)^{2}}{n^{2}}\left(\log \left(1-F^{*}\left(y_{i}\right)\right)-\log \left(1-F^{*}\left(y_{i+1}\right)\right)-1\right)\right.\right. \\
& =18\left(\frac{145.228528}{18^{2}}+\frac{223.261928}{18^{2}}-1\right) \\
& =2.47169199996
\end{aligned}
$$

5. An insurance company collects a sample of 1500 claims. They want to decide whether this data is better modeled as following an inverse exponential distribution, or a generalised Pareto distribution. After calculating MLE estimates for the parameters (1 parameter for the inverse exponential and 3 for the generalised Pareto), log-likelihoods for the two distributions are:

| Distribution | log-likelihood |
| :--- | :--- |
| Inverse Exponential | -4244.75 |
| Generalised Pareto | -4236.89 |

Use a BIC to decide whether the generalised Pareto distribution or the inverse exponential distribution is a better fit for the data.

The BIC for the inverse exponential is $-4244.75-\frac{1}{2} \log (1500)=-4248.40661019$. The BIC for the Generalised Pareto is $-4236.89-\frac{3}{2} \log (1500)=-4247.85983058$. Therefore, the Generalised Pareto is prefered.
6. A homeowner's house has a value of $\$ 860,000$. The insurer requires $75 \%$ coverage for full insurance. The deductible is $\$ 4,000$, decreasing linearly to zero for losses of $\$ 12,000$. The home sustains $\$ 7,000$ of damage from fire. The insurer reimburses $\$ 3,300$. For what value was the home insured?

The deductible for a loss of $\$ 7,000$ is $\frac{12000-7000}{12000-4000} \times 4000=\$ 2,500$. Thus, under, full insurance, the insurer would reimburse $7000-2500=\$ 4,500$. Therefore, the coverage for this home is $\frac{3300}{4500}=\frac{11}{15}$. For full insurance, the home should be insured for $860000 \times 0.75=\$ 645,000$. Therefore, the home is insured for $\frac{11}{15} \times 645000=\$ 473,000$.
7. The following table shows the cumulative losses (in thousands) on claims from one line of business of an insurance company over the past 4 years.

|  | Development year |  |  |  |
| ---: | ---: | ---: | ---: | ---: |
| Accident year | 0 | 1 | 2 | 3 |
| 2018 | 5539 | 6003 | 6829 | 7108 |
| 2019 | 6243 | 6792 | 7314 |  |
| 2020 | 6217 | 7209 |  |  |
| 2021 | 6372 |  |  |  |

Using the mean for calculating loss development factors, esimate the total reserve needed for payments to be made in 2023 (two years in the future) using the Bornhuetter-Fergusson method. The expected loss ratio is 0.74 and the earned premiums in each year are given in the following table:

| Year | Earned <br> Premiums (000's) |
| :---: | :---: |
| 2018 | 8537 |
| 2019 | 8764 |
| 2020 | 9023 |
| 2021 | 9285 |

[Assume no more payments are made after development year 3.]
The mean loss development factors are

| Development Year | Loss Development Factor |
| :--- | :--- |
| $0 / 1$ | $\frac{20004}{17999}=1.1113950775$ |
| $1 / 2$ | $\frac{1413}{12795}=1.10535365377$ |
| $2 / 3$ | $\frac{7108}{6829}=1.04085517645$ |

The proportion of cumulative payments by the end of each development year are:

| Development Year | Cumulative Proportion <br> of payments made | Proportion of <br> payments made |
| :--- | ---: | ---: | ---: |
| 0 | $\overline{1.1113950775 \times 1.10535365377 \times 1.04085517645}=0.782059819772$ | 0.782059819772 |
| 1 | $\overline{1.10535365377 \times 1.04085517645}=0.869177434008$ | 0.087117614236 |
| 2 | $\frac{1}{1.04085517645}=0.960748452451$ | 0.091571018443 |
| 3 | 1 | 0.039251547549 |

The expected payments to be made in 2023 from Accident year 2020 are $9023 \times 0.74 \times 0.039251547549=262.083368016$. The expected payments to be made in 2023 from Accident year 2021 are $9285 \times 0.74 \times 0.091571018443=629.17531062$. Thus, the total reserves needed for payments to be made in 2023 are $262.083368016+629.17531062=\$ 891.26$.

