

## ACSC/STAT 3703 - Midterm 2 Solutions - out of 30 points

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### Formulas:

- If  $X$  has an exponential distribution with mean  $\theta$ , then the distribution function of  $X$  is  $F_X(x) = 1 - e^{-x/\theta}$ , and  $E[X \wedge c] = \theta F_X(c)$ .
- The p.g.f of a Poisson random variable with mean  $\lambda$  is  $P_X(z) = e^{\lambda(z-1)}$ .
- For any pair of random variables  $X$  and  $Y$ ,  $E[Y] = E[E[Y|X]]$  and  $V[Y] = E[V[Y|X]] + V[E[Y|X]]$ .

1. The number of claims  $N$  made to an insurance company in a given year is a random variable with mean 100 and standard deviation 10. The individual claims  $X_1, X_2, \dots$  are independent of one another and of  $N$ . The mean claim size is  $\mu_X = 500$  and the standard deviation of the claim size is  $\sigma_X = 10$ . The total of the claims made against the company in one year is  $S_N = X_1 + X_2 + \dots + X_N$ . (Hint: for both parts of this question, use the mean and variance formulas above.)

(3) (a) What is the mean of  $S_N$ ?

$$E[S] = E[X]E[N] = 500(100) = 50,000$$

(b) What is the standard deviation of  $S_N$ ?

$$V[S] = E[V[S|N]] + V[E[S|N]] = E[NV(X)] + V(NE[X])$$

$$= E[N]V(X) + (E[X])^2V[N] = 100(100) + 500^2(100) = 25,010,000$$

(3) standard deviation is  $\sqrt{25,010,000} \approx 5001$ .

2. Suppose that  $X$  is a discrete random variable with probability mass function given by  $P(X = k) = p_k, k = 0, 1, 2, \dots$ . Let  $X^T$  be the 0 truncated version of  $X$ , for which

$$P(X^T = k) = p_k^T = \frac{p_k}{1 - p_0}, k = 1, 2, \dots$$

and  $X^M$  be the 0 modified version of  $X$ , for which

$$P(X^M = k) = (1 - p_0^M)p_k^T, k = 1, 2, \dots$$

(a) Prove that  $E(X^M) = E(X) \frac{1-p_0^M}{1-p_0}$ .

$$P_k^M = (1 - p_0^M)p_k^T = \frac{1-p_0^M}{1-p_0}p_k, \text{ for } k=1,2, \dots$$

(3) The result follows immediately.

(b) The following table gives the probability distribution of a random variable  $X$ . Fill in the missing entries for the probability distributions of the zero truncated and zero modified random variables.

$$\frac{1 - p_0^M}{1 - p_0} = \frac{p_2^M}{p_2} = (1/6)/(1/4) = 2/3$$

so  $p_k^M = p_k(2/3)$ , for  $k=1,2,3$ . Then  $p_0^M = 1 - p_1^M - p_2^M - p_3^M$ .  
for  $k = 1, 2, 3, p_k^T = p_k/(1 - p_0) = p_k(4/3)$ , giving the following

(4)

$k$	$P(X = k)$	$P(X^T = k)$	$P(X^M = k)$
0	1/4	0	1/2
1	3/8	1/2	1/4
2	1/4	1/3	1/6
3	1/8	1/6	1/12

3. Losses follow an exponential distribution with mean  $\lambda$ . (Recall that the loss elimination ratio associated with an ordinary deductible  $d$  is  $\frac{E(X \wedge d)}{E(X)}$ . What is the size of the ordinary deductible needed to give a loss elimination ratio of  $1/2$ ?

$$\frac{E(X \wedge d)}{E(X)} = 1 - e^{-d/\lambda} = 1/2 \rightarrow d = -\lambda \log_e(1/2) \approx .693\lambda$$

(5)

4. From theorem 6.1 we know that if  $X_1, X_2, \dots, X_n$  are i.i.d. Poisson with mean  $\lambda$  then  $S_n = \sum_{i=1}^n X_i$  is Poisson with mean  $n\lambda$ . The theorem was proven by finding the probability generating function of  $S_n$ , beginning with the common pgf of the  $X$ 's.

Now consider  $S_N = \sum_{i=1}^N X_i$ , where  $N$  is independent of  $X_1, X_2, \dots$  and the  $X$ 's are i.i.d. The p.g.f. of  $S_N$  is derived in section 9.3.1 and is given by  $P_{S_N}(z) = P_N(P_X(z))$  where  $P_N(z) = E[z^N]$  is the p.g.f. of  $N$  and  $P_X(z) = E[z^X]$  is the common p.g.f. of the  $X$ 's.

Suppose that the  $X_j$ 's are i.i.d. Poisson( $\lambda$ ), independent of  $N$ , and  $N$  is Poisson with mean  $\gamma$ .

- (a) What is the p.g.f. of  $S_N$ ?

$$P_S(z) = P_N(P_X(z)) = P_N(e^{\lambda(z-1)}) = e^{\gamma(e^{\lambda(z-1)} - 1)}$$

(3)

- (b) Does  $S_N$  have a Poisson distribution? Why?

No, because the above pgf of  $S$  cannot be written as  $e^{\theta(z-1)}$  for any  $\theta$ .

(2)

5. The amount of a loss  $X$  has density function

$$f_X(x) = \frac{1}{100}, \quad 0 < x < 100$$

Insurance pays 80% of the amount of a loss in excess of an ordinary deductible of  $d = 20$ . The maximum insured loss is  $u = 60$ .

(a) What is  $E[Y_L]$ , the expected value of the loss?

Use the result of theorem 8.7, which states that

$$E(Y^L) = \alpha(1+r) \left[ E\left(X \wedge \frac{u}{1+r}\right) - E\left(X \wedge \frac{d}{1+r}\right) \right]$$

(5) where  $r$  is the inflation rate, 0 in this case, and  $\alpha$  is the co-payment proportion - the proportion paid by the insurance company.

$$\begin{aligned} E[X \wedge c] &= \int_{x=0}^c x f(x) dx + c \int_{x=c}^{100} f(x) dx \\ &= c^2/200 + c(100 - c)/100 = c - c^2/200 \end{aligned}$$

so

$$E(Y^L) = .8(60 - 60^2/200 - 20 + 20^2/200) = .8(60 - 18 - 20 + 2) = 24(.8) = 19.2$$

(b) What is the expected value of the payment  $E[Y_P]$ ? (Hint: Recall that a payment is made only if  $X > 20$ .)

$$E[Y^P] = \frac{E[Y^L]}{P(X > 20)} = 19.2/.8 = 24$$

(2)

(answer to part a divided by .8)